JUST EE Dept.

EE301 Applied Probability Theory Final Exam

August 17th, 1997

The exam has 42 questions, for a total of 42 points. Answer All Questions

- 1. The largest absolute value of the joint characteristic function for N random variables
 - (a) is bound by the largest value of the joint pdf
 - (b) =1
 - (c) is bound by the largest values of its marginal pdfs
 - (d) =0
 - (e) = the value of the pdf at the origin $(e) = e^{-i\theta}$
- 2. Three uncorrelated r.v.s X_1, X_2 and X_3 have means $\bar{X}_1 = 1, \bar{X}_2 = -3$, and $\bar{X}_3 = 1.5$ and second moments $E[X_1^2] = 2.5, E[X_2^2] = 11$, and $E[X_3^2] = 3.5$. Let $Y = X_1 - 2X_2 + 3X_3$ be a new r.v., its mean=
 - (a) -0.5
 - (b) 0
 - (c) 0.5
 - (d) 2.0
 - (e) 11.5
- 3. In the above question, the variance of Y =
 - (a) 4.75
 - (b) 9.25
 - (c) 9.5
 - (d) 20.75
 - (e) 41.5
- 4. Three statistically independent random variables X, Y, and Z having means and variances as follows: $\bar{X} = -1, \sigma_X^2 = 2.0, \bar{Y} = 0.6, \sigma_Y^2 = 1.5, \bar{Z} = 1.8, \sigma_Z^2 = 0.8$. Using central limit theorem, the r.v. W = X + Y + Z have a gaussian statistics of mean and variance equal to
 - (a) 1.4 and 8.6
 - (b) 2.4 and 8.6
 - (c) $\frac{1.4}{3}$ and $\frac{4.3}{3}$
 - (d) 1.4 and 4.3
 - (e) 2.4 and 4.3

5. Two random variables X and Y have a joint pdf

$$f_{X,Y}(x,y) = \begin{cases} \frac{5}{16}x^2y, & 0 < y < x < 2\\ 0, & elsewhere \end{cases}$$

X and Y

- (a) are statistically independent
- (b) are not statistically independent
- (c) have same mean
- (d) have same variance
- (e) have the same second moment
- 6. Given the function

$$f_{X,Y}(x,y) = \begin{cases} (x^2 + y^2)/(8\pi), & x^2 + y^2 < 4\\ 0, & elsewhere \end{cases}$$

Then $P\{2 < X^2 + Y^2 \le 3.2\} =$

- (a) 0.12
- (b) 0.39
- (c) 0.5
- (d) 0.61
- (e) 0.88
- 7. The function

$$G_{X,Y}(x,y) = \begin{cases} 0, & x < y \\ 1, & x \ge y \end{cases}$$

is a valid distribution function.

- (a) True
- (b) False
- 8. If X and Y are independent, then for all y we have E[X|Y = y] =
 - (a) E[X] E[Y]
 - (b) E[Y|X = x]
 - (c) \bar{Y}
 - (d) E[X]
 - (e) $f_X(X|Y=y)$

- 9. If Y = a + bX, then $\rho(X, Y) = 1$ if
 - (a) b > 1
 - (b) b > 0
 - (c) a = b
 - (d) b < 0
 - (e) a = -b
- 10. If X and Y are identically distributed, not necessarily independent, then Cov(X+Y, X-Y) =
 - (a) Cov(X, X Y)
 - (b) $\sigma_X^2 + \sigma_Y^2$
 - (c) Cov(X+Y,X)
 - (d) 0
 - (e) Cov(Y, X Y)
- 11. The power density spectrum of a r.p. X(t) is 16. When a fair coin is tossed, the possible outcomes, given by $S_{XX}(\omega) = Ae^{-\omega^2/(2\omega_o^2)}$. Its autocorrelation function is given by
 - (a) $R_{XX}(\tau) = \frac{A\omega_o}{\sqrt{2\pi}} e^{-(\omega_o \tau^2)/2}$

(b)
$$R_{XX}(\tau) = \frac{A\omega_o}{\sqrt{2\pi}}e^{-(\omega_o\tau)/2}$$

(c)
$$R_{XX}(t, t+\tau) = \frac{A\omega_o}{\sqrt{2\pi}} e^{-(\omega_o(t-\tau)^2)/2}$$

(d)
$$R_{XX}(\tau) = \frac{A\omega_o}{\sqrt{2\pi}} e^{-(\omega_o^2 \tau^2)/2}$$

(e) $R_{XX}(\tau) = \frac{A\omega_o}{\sqrt{2\pi}} (1 - (\omega_o \tau^2)^2)$

- 12. The power density spectrum of a r.p. X(t) is 17. Find k for the joint pdf given by $S_{XX}(\omega) = A$ for $-\omega_o < \omega < \omega_o$ and zero elsewhere. Its autocorrelation function is given by $R_{XX}(\tau) = \frac{A}{\pi\tau} sin(\omega_o \tau)$
 - (a) true
 - (b) false
- 13. If X(t) and Y(t) are two independent random processes with power density spectra $S_{XX}(\omega)$ and $S_{YY}(\omega)$, the power density spectrum of their linear combination Z(t) = aX(t) + bY(t) is

(a)
$$S_{ZZ}(\omega) = S_{XX}(\omega) + S_{YY}(\omega)$$

(b) $S_{ZZ}(\omega) = 2ab(S_{XX}(\omega) + S_{YY}(\omega))$
(c) $S_{ZZ}(\omega) = a^2 S_{XX}(\omega) - b^2 S_{YY}(\omega)$
(d) $S_{ZZ}(\omega) = a S_{XX}(\omega) + b S_{YY}(\omega)$
(e) $S_{ZZ}(\omega) = a^2 S_{XX}(\omega) + b^2 S_{YY}(\omega)$

14. The output Y(t) of a linear system is k times the input X(t). Then

- (a) $R_{YY} = k^2 R_{XX}$ (b) $R_{YY} = kR_{XX}$
- (c) $R_{YY} = R_{XX}$
- (d) $R_{YY} = R_{XX}/k$
- (e) $R_{YY} = R_{XX}/k^2$
- 15. The characteristic function for a r.v. X having exponential pdf is given by $\Phi_X(\omega) = \frac{1}{1-ia\omega}$, then its variance is
 - (a) $-a^2$
 - (b) -a
 - (c) ja^2
 - (d) *a*
 - (e) a^2
 - heads and tails, are represented by a random variable X which has the value 1 or -1, respectively. The characteristic function equals
 - (a) $(1 + e^{j\omega})/2$ (b) $j\sin(\omega)$ (c) $-j\sin(\omega)$ (d) $-\cos(\omega)$ (e) $\cos(\omega)$

$$f_{X,Y}(x,y) = \begin{cases} ke^{-2x-3y}, & x, y \ge 0\\ 0 & else \end{cases}$$
(a) 6
(b) 1
(c) 1/2
(d) 1/3

(e) 1/6

18. Find μ_{11} for the joint pdf

$$f_{X,Y}(x,y) = \begin{cases} ke^{-(x+y)}, & x, y \ge 0\\ 0 & else \end{cases}$$
(a) -1
(b) 0
(c) 1/2
(d) 1
(e) ∞

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- 19. A random variable x is uniformly distributed between -1 and +1. Determine the normalized correlation coefficient for x and y if $y = x^2$.
 - (a) -1
 - (b) -8/12
 - (c) 0
 - (d) 1
 - (e) 4/12
- 20. A random variable X has a probability density function $f_X(x) = \begin{cases} e^{-x}, & x \ge 0 \\ 0 & else \end{cases}$ Determine the probability density function for the random variable $Y = X^3$.
 - (a) $e^{-y}/3y^2$ (b) $e^{-y^{1/3}}/3y^{2/3}$ (c) $2e^{-y}/y$ (d) $e^{-y^3}/3y^2$
 - (e) $e^{y^{1/3}}/3y^{-2/3}$
- 21. The magnitude of the characteristic function of a random variable X is
 - (a) ≥ 1.0
 - (b) $\geq \bar{X}$
 - (c) ≤ 1.0
 - (d) $\leq \sigma_X$
 - (e) $\leq \sigma_X^2$
- 22. Chebyshev's inequality states that any random variable X has a finite second moment, then for any constant C > 0 we have
 - (a) $P(|X| \ge c) \le \frac{E(X^2)}{c^2}$ (b) $P(|X| \le c) \le \frac{E(X^2)}{c^2}$ (c) $P(|X| \ge c) \le c^2 E(X^2)$ (d) $P(|X| \ge c) \le \frac{E(X)}{c^2}$ (e) $P(|X| \le c) \le c^2 E(X^2)$
- 23. An analogue of Chebyshev's inequality where the absolute first moment is used in place of the second moment can be written as P(|X m|) > c)

(a)
$$\geq E(|X-m|)$$

(b)
$$\leq E(|X - m|)$$

- (c) $\leq E(|X|)$ (d) $\leq E(X^2)$ (e) $\geq E(|X|)$
- 24. For two independent random variable X and Y, with known individual pdf's, then any new random variable formed directly from their sum; i.e., S = X + Y; has a pdf equal to
 - (a) $f_S(s) = f_X(x)f_Y(y)$
 - (b) $f_S(s) = \int_{-\infty}^{\infty} f_Y(y) f_X(s-y) dy$

(c)
$$f_S(s) = f_X(x) + f_Y(y)$$

- (d) $f_S(s) = \frac{f_X(x)}{f_Y(y)}$ (e) $f_S(s) = \int_{-\infty}^{\infty} f_X(x) f_Y(s-y) dx$
- 25. A geometric random variable X has the probability $P(X = k) = pq^k, k = 0, 1, 2, \dots, p + q = 1$, its expected value is
 - (a) $\frac{1}{p}$
 - (b) *pq*
 - (c) $\frac{p}{q}$
 - (d) kpq
 - (e) $\frac{q}{p}$
- 26. A r.v. X drawn from a Pascal distribution (or negative binomial),

$$P[X=k] = \left(\begin{array}{c} n+k-1\\k \end{array}\right) p^n q^k$$

, $k = 0, 1, \dots, p + q = 1$. Its mean value=

- (a) nq/p
- (b) q/p
- (c) p/q
- (d) np/q
- (e) kp/q
- 27. For any r.v. X with characteristic function $\Phi(\omega) = E[e^{j\omega X}]$, then for any collection of n complex numbers a_i we have $E[|\sum_{i=1}^n a_i e^{j\omega_i X}|^2] \ge 0$, one can write $\sum_{i=1}^n \sum_{k=1}^n \Phi(\omega_i \omega_k) a_i a_k^* \ge 0$
 - (a) true
 - (b) false

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28.

$$f_{X,Y}(x,y) = \begin{cases} \frac{15}{2}x(2-x-y) & 0 < x, y < 1\\ 0 & \text{otherwise} \end{cases}$$

then the conditional pdf of X, given that Y = y equals to

(a)
$$\frac{6(2-x-y)}{4-3y}$$

(b) $\frac{6x(2-x-y)}{4-3y}$
(c) $\frac{x(2-x-y)}{6(4-3y)}$
(d) $\frac{6x^2(2-x-y)}{3y}$
(e) $\frac{x(2-x-y)}{4}$

29. The indicator variable I for an event A is defined as

$$I = \begin{cases} 1, & \text{if } A \text{ occurs} \\ 0, & \text{if } \bar{A} \text{ occurs} \end{cases}$$

We can write

(a)
$$P(I) = E[A]$$

(b)
$$E[I] = P(A)$$

(c)
$$Var[I] = E[A]$$

- (d) Var[A] = E[I]
- (e) $P(\bar{A}) = E[I]$
- 30. In a group of resistors where 4% are of metal type resistors, what is the probability that there are no metal resistors in a random selection of 25 resistors?
 - (a) $1.0 (0.04)^{25}$
 - (b) $1 (0.96)^{25}$

(c)
$$(0.04)^{25}$$

- (d) 0.96
- (e) $(1 0.04)^{25}$
- 31. Given that

$$f_{X|Y}(X|Y) = \begin{cases} ye^{-xy} & 0 < x < \infty, 0 < y < 2\\ 0 & \text{otherwise} \end{cases}$$

then $E[e^{x/3}|Y=1]$ is

(c) 3/2

- (d) 2
- (e) 2/3

32. Let X be a r.v. and c be a real constant number, let Y = X + c, then $\sigma_Y^2 =$

- (a) $c^2 \sigma_X^2$ (b) $\sigma_X^2 + c$ (c) $\sigma_X^2 + c^2$ (d) σ_X^2 (e) $\sigma_X^2 - c^2$
- 33. If two points X and Y are picked at random from the interval [0,1], the probability that the distance between them $P(|X - Y| < \frac{1}{2}) =$
 - (a) 1/8
 - (b) 1/4
 - (c) 3/8
 - (d) 1/2
 - (e) 3/4
- 34. Let X_1, X_2, \dots, X_n be independent random variables with cdfs F_1, F_2, \cdots, F_n . $= maximum(X_1, X_2, \cdots, X_n)$ and Let M $minimum(X_1, X_2, \cdots, X_n).$ m=The distribution for M can be written as $F_{max}(x) = P(M \leq x) = P(X_1 \leq x; X_2 \leq$ $x; \cdots; X_n \leq x) = P(X_1 \leq x)P(X_2 \leq x)$ $x)\cdots P(X_n \leq x) = F_1(x)F_2(x)\cdots F_n(x).$ Following the arguments mentioned one can write the cdf for m, the minimum, as
 - (a) $F_{min}(x) = \prod_{i=1}^{n} (1 F_i(x))$
 - (b) $F_{min}(x) = 1 \prod_{i=1}^{n} F_i(x)$
 - (c) $F_{min}(x) = \prod_{i=1}^{n} F_i(x)$
 - (d) $F_{min}(x) = 1 \sum_{i=1}^{n} (1 F_i(x))$
 - (e) $F_{min}(x) = 1 \prod_{i=1}^{n} (1 F_i(x))$
- 35. Ergodic process X(t) is a stochastic process that
 - (a) is second-order stationary
 - (b) has statistical averages = time averages
 - (c) is strictly stationary
 - (d) is first-order stationary
 - (e) has $\bar{X} = \text{constant}$, and $R_{XX}(\tau) = \text{constant}$.

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36. If
$$X(t)$$
 is WSS, and $Z(t) = aX(t)$ then
(a) $C_{ZZ}(\tau) = |a|^2 C_{XX}(\tau)$
(b) $R_{XZ}(\tau) = |a|^2 R_{XX}(\tau) R_{ZZ}(\tau)$
(c) $R_{ZX}(\tau) = |a|^2 R_{XX}(\tau) R_{ZZ}(\tau)$
(d) $R_{ZZ}(\tau) = |a|^2 R_{XX}(\tau)$
(e) $C_{XZ}(\tau) = |a|^2 C_{ZX}(\tau)$

- 37. A r.v. X drawn from a binomial distribution, i.e. $P(X = k) = \binom{N}{k} p^k (1-p)^{N-k}$ then its variance = (a) Np(1-p)(b) Np(c) N(1-p)(d) (N-1)p(1-p)(e) $N^2p(1-p)$
- 38. R.V. X has

$$f_X(x) = \begin{cases} p & x = 1\\ 1 - p & x = -1 \end{cases}$$

Its characteristic function =

- (a) $p + (1-p)e^{j\omega}$ (b) $e^{-j\omega} + 2jp\sin(\omega)$ (c) $(1-p)e^{j\omega} + pe^{-j\omega}$ (d) $(1-p)e^{-j\omega}$
- (e) $pe^{j\omega}$
- 39. For the Poisson distribution, $P_X(X = k) = \frac{e^{-\lambda}\lambda^k}{k!}$ its mean = λ and its variance =
 - (a) λ^2
 - (b) $\frac{1}{\lambda}$
 - (c) $k\lambda$
 - (d) $e^{-\lambda}$
 - (e) λ
- 40. Random Walk problem: Define $S = \sum_{i=1}^{N} X_i$ and $X_i, i = 1, \dots, N$ are two-value i.i.d. (independent identically distributed) random variables, $P(X_i = +1) = p$ and $P(X_i = -1) = 1 p = q$ then E[S] =
 - (a) N

- (b) $N \times p$ (c) $N \times q$ (d) N(p-q)(e) 0
- 41. $S_{XX}(\omega) = \int_{-\infty}^{\infty} R_{XX}(\tau) e^{-j\omega\tau} d\tau$ is valid if X(t) is
 - (a) first-order stationary
 - (b) second-order stationary
 - (c) any process with zero mean
 - (d) ergodic process
 - (e) any process with unit variance
- 42. Telephone calls are initiated through an exchange at the average rate of 30 per minute and are described by a Poisson process. The probability that less than 3 calls are initiated in any 2-second period =
 - (a) $1 e^{-b}[1 + b + b^2/2]$ with b = 1
 - (b) $1 e^{-b}[1+b]$ with b = 0.5
 - (c) $e^{-b}[1+b+b^2/2]$ with b=1
 - (d) $e^{-b}[1+b]$ with b = 2.0
 - (e) $1 e^{-b}[1+b]$ with b = 1